

**Acknowledgements** This work was supported by NSF Chemical and Biological Oceanography and by NOAA Ocean Global Carbon Cycle Program. We are grateful to T. Gunderson and Y. Zhang for technical support.

**Competing interests statement** The authors declare that they have no competing financial interests.

**Correspondence** and requests for materials should be addressed to S.A.S.W. (ssanudo@notes.cc.sunysb.edu).

## Public health vaccination policies for containing an anthrax outbreak

Ron Brookmeyer<sup>1</sup>, Elizabeth Johnson<sup>1</sup> & Robert Bollinger<sup>2</sup>

<sup>1</sup>Department of Biostatistics, Johns Hopkins Bloomberg School of Public Health, 615 North Wolfe Street, Baltimore, Maryland 21205, USA

<sup>2</sup>Division of Infectious Diseases, Johns Hopkins University School of Medicine, 720 Rutland Avenue, Baltimore, Maryland 21205, USA

Concern about biological weapons has raised questions about the most effective public health policies to contain an anthrax outbreak<sup>1–3</sup>. We developed a probability model to predict the impact of different anthrax antibiotic and vaccination policies. An anthrax outbreak can be significantly contained by minimizing the delay until initiation of antibiotic prophylaxis. However, even if mass distribution of antibiotics is completed within six days of the initial exposure, then at most about 70% of cases can be prevented. Post-exposure vaccination will not significantly increase that prevention rate if adherence to antibiotic regimens is similar or higher than that attained in the 2001 US outbreak<sup>4</sup>. However, post-exposure vaccination can be useful either in shortening the duration of a prolonged antibiotic regimen, in the event of an antibiotic-resistant strain, or if antibiotic adherence rates are very low. Here we show that a mass pre-exposure vaccination programme for the general population would require very high population coverage rates to significantly increase prevention rates from that achieved with targeted and rapid post-exposure prophylaxis programmes.

Public health planners are uncertain how an anthrax vaccine should be used as part of a comprehensive strategy for addressing the bioterrorism threats posed by *Bacillus anthracis*<sup>1–3</sup>. Critical public health questions include whether an anthrax vaccine should be used before exposure to immunize the general population, or used as a post-exposure prophylaxis. The answers depend on many factors including the safety and efficacy of prolonged antibiotic prophylaxis, adherence to antibiotic regimens, the time delay before post-exposure antibiotic prophylaxis is initiated and vaccine characteristics such as efficacy, safety and time to achieve immunity. The US Food and Drug administration licensed an anthrax vaccine, Anthrax Vaccine Absorbed (AVA) in 1970. Because AVA can require six doses and up to 18 months to achieve maximum protection, research has accelerated to develop and manufacture an improved anthrax vaccine. It is hoped that recombinant protective antigen (rPA) anthrax vaccine could provide maximum protection in three doses over a period of not more than 28 days<sup>2,5–6</sup>.

The Centers for Disease Control and Prevention recommended that those people exposed in the 2001 US anthrax attacks receive 60 days of antibiotic prophylaxis. Subsequently, it was recommended that people consider extending the antibiotic therapy with or without three doses of the anthrax vaccine<sup>3,7</sup>. However, few opted to take this additional recommendation, and adherence to the initially recommended 60-day antibiotic regimen was less than 50%<sup>8</sup>. Nevertheless, no persons among the more than 10,000

recommended for prophylaxis therapy developed disease and among them only a handful of cases were prevented<sup>9</sup>. This experience has raised questions about which public health strategies most effectively contain an anthrax outbreak.

A 1-kg release of an anthrax preparation could contain more than 10<sup>14</sup> spores, roughly equivalent to five-billion lethal doses<sup>10–11</sup>. Although the magnitude of such a release seems daunting, the critical determinants for public health involve not just the size of the release but also dispersal factors, which include the size of the spores, the type of spore preparation (wet or dry), method of release, wind conditions, temperatures, building insulation and the population density of the exposed areas (both inside and outside of buildings)<sup>2,10–12</sup>. Atmospheric models predict that a 100-kg release could result in between tens of thousands to several million deaths, depending on these factors<sup>2,11</sup>. In the face of this uncertainty, how does one determine rational public health policy?

Quantitative models have been used recently to inform the policy debate about smallpox vaccination strategies<sup>13–17</sup>. To evaluate strategies for containing an anthrax outbreak, we developed a probability model that accounts for the dynamics of spore clearance and germination, and describes the effects of pre-exposure vaccination, post-exposure vaccination, antibiotic prophylaxis and antibiotic adherence. We present the results in terms of the percentages of cases that can be prevented with various public health policies. We show that these percentages do not vary appreciably by exposure level (dose of inhaled spores), thus allowing public health policy to be guided, even with the great uncertainty surrounding the specifics of bioterrorism threats.

We considered three models for the dose of inhaled spores: a fixed high-exposure level in which all persons are exposed to the infectious dose ID(50), which yields a rate of 5,000 cases per 10,000

Table 1 Per cent of cases of anthrax prevented by post-exposure intervention

Antibiotic response time*	Antibiotic duration (days)	Antibiotic adherence†	Percentage of cases prevented Post-exposure vaccination‡	
			No	Yes
High exposure§				
None	None	None	0	6
Slow	60	Partial	40	44
Slow	60	Full	46	47
Slow	120	Full	47	47
Rapid	60	Partial	58	64
Rapid	60	Full	68	69
Rapid	120	Full	69	69
Variable moderate exposure§				
None	None	None	0	6
Slow	60	Partial	39	43
Slow	60	Full	46	46
Slow	120	Full	46	46
Rapid	60	Partial	57	63
Rapid	60	Full	67	68
Rapid	120	Full	68	68
Low exposure§				
None	None	None	0	9
Slow	60	Partial	48	52
Slow	60	Full	55	55
Slow	120	Full	55	55
Rapid	60	Partial	66	71
Rapid	60	Full	75	76
Rapid	120	Full	76	76

\*Efficacy of antibiotics is assumed to be 100% (that is, the probability of disease protection if on antibiotics at the time of spore germination is 1). With slow response, mass antibiotic distribution begins 6 days after initial exposure and the distribution is completed over the following 6 days (total response time = 12 days). With rapid response, mass antibiotic distribution begins 3 days after initial exposure and distribution is completed over the following 3 days (total response time = 6 days).

†With full adherence, 100% of persons complete the recommended duration of antibiotics (60 or 120 days). With partial adherence to a 60-day regimen, 25% of persons each complete 15, 30, 45 and 60 days of antibiotics.

‡Post-exposure vaccine protection (with 95% efficacy) is assumed to begin 28 days after persons begin antibiotics. In the case of no antibiotics, post-exposure vaccine protection begins 28 days after the sixth day following exposure.

§High exposure is the ID(50). Variable moderate exposure assumes that the percentages exposed to the ID(90), ID(50), ID(10) and ID(1) are respectively 2%, 10%, 28% and 60%, with an average about the ID(10) of 1,020 cases out of 10,000 exposed, if no intervention. Low exposure is the ID(1).

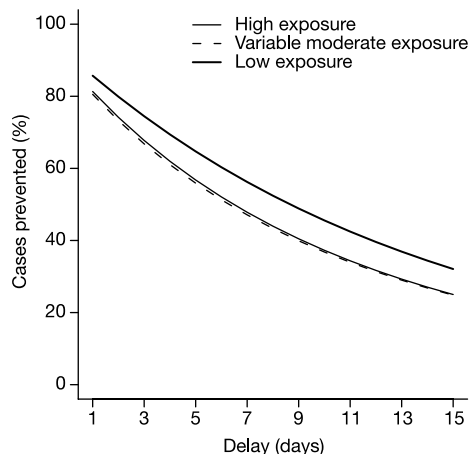
exposed persons with no public health intervention; a fixed low-exposure level in which all persons are exposed to the ID(1) and which yields a rate of 100 cases per 10,000 exposed; and a variable moderate exposure level in which the numbers of inhaled spores varied across the population and yields an average rate of 1,020 cases per 10,000 exposed.

Table 1 displays the percentage of cases prevented with various post-exposure interventions. Surprisingly, the results do not vary appreciably by exposure level. For example, with a rapid 6-day response in which mass antibiotic distribution begins 3 days after the initial exposure and is completed after an additional 3 days, the percentage of cases prevented with full 60-day antibiotic adherence were 68%, 67% and 75% for the high, variable moderate and low exposures, respectively. With a slower 12-day response in which mass antibiotic distribution begins 6 days after the initial exposure and is completed after an additional 6 days, the corresponding percentages of cases prevented were 46%, 46% and 55%. Figure 1 shows the relationship between the percentage of cases prevented and the delay until beginning mass antibiotic distribution. The figure shows that prevention rates, which do not vary significantly by exposure level, are high with short delays, but fall to under 50% after a 10-day delay.

We considered the situation when there was only partial adherence to the antibiotic regimen; specifically, when 25% of people completed either 15, 30, 45 or 60 days of a 60-day regimen, similar to that observed in the 2001 US anthrax outbreak<sup>8</sup>. The percentages of cases prevented drops by about 7–10% compared with full antibiotic adherence (Table 1).

We evaluated the preventive value of post-exposure vaccination as an adjuvant to antibiotic prophylaxis. Our first set of calculations assumed a 95% vaccine efficacy with a 28-day delay to achieve vaccine-induced immunity<sup>5</sup>. If people adhere to the antibiotic regimen, then addition of post-exposure vaccination increases the per cent prevented by, at most, 1% (Table 1). If people only partially adhere to the antibiotic regimen, post-exposure vaccination modestly increases the prevention rate by no more than 6% (Table 1). If the strain of *B. anthracis* is resistant to antibiotics, then post-exposure vaccination prevents between 6% and 9% of cases depending on the exposure level (row 1, Table 1). These modest preventive effects of post-exposure vaccination are explained by the time delay to achieve vaccine-induced immunity.

We performed a sensitivity analysis to the time delay to achieve vaccine-induced immunity, to identify characteristics that would improve a vaccine's utility as a post-exposure prophylaxis. Table 2



**Figure 1** Percentage of cases prevented versus delay of initiating mass distribution of antibiotics. Calculations based on completing mass distribution of antibiotics over 3 days, and assuming that persons adhere completely to a 60-day regimen and that antibiotic efficacy is 100%. Exposure levels are defined in Table 1.

**Table 2 Sensitivity of prevention rates to post-exposure vaccine immunogenicity**

Cumulative vaccine protection* (%)					Percentage of cases prevented Antibiotic response†						
					No adherence 0% efficacy	Speed: slow		Speed: slow		Speed: rapid	
						Partial adhere		Partial adhere		Partial adhere	
						90% efficacy		100% efficacy		100% efficacy	
7	14	21	28								
High exposure‡											
0	0	0	0		0	36		40		58	
0	0	0	95		6	40		44		64	
0	0	25	95		7	41		44		65	
0	25	50	95		10	42		45		66	
25	50	75	95		17	43		46		67	
Low exposure‡											
0	0	0	0		0	43		48		66	
0	0	0	95		9	48		52		71	
0	0	25	95		10	48		53		72	
0	25	50	95		14	49		53		73	
25	50	75	95		22	51		54		74	

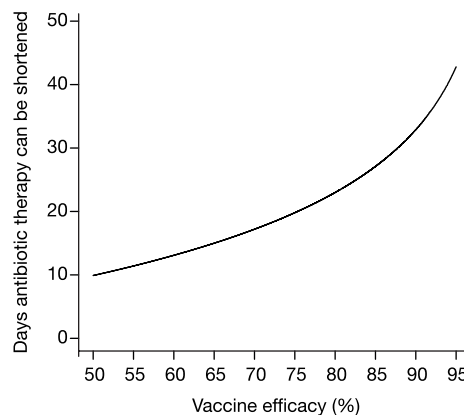
\* Cumulative per cent of persons fully protected by vaccine following days 7, 14, 21 and 28 after receiving the first dose of vaccine. As described in footnotes of Table 1, first vaccine dose is given when antibiotics are initiated, or in the case of no antibiotics on day 6 following exposure.

† Antibiotic efficacy (% of persons in whom antibiotics prevented disease) is 0% (as in case of antibiotic-resistant strain), 90% or 100%. Speed in distributing antibiotics and adherence rates are described in footnotes of Table 1.

‡ See footnotes to Table 1.

shows the percentages of cases prevented under different assumptions of vaccine immunogenicity. For example, we considered a highly immunogenic vaccine in which 25%, 50%, 75% and 95% of people achieve immunity within 7, 14, 21 and 28 days following the first vaccine dose, respectively (last row, Table 2). We find that the addition of such a vaccine to an effective antibiotic regimen results in an increase in the prevention rate of no more than 9% in all situations considered (including one in which antibiotics were only partially (90%) effective). A vaccine that also reduced disease severity could have an additional benefit of increasing the percentages of deaths prevented. For example, suppose people are exposed to high doses of an antibiotic-resistant strain; if a single dose of a highly immunogenic vaccine given before disease onset also reduced the mortality rate among cases by 50%, 25%, 10% and 0%, then the percentages of deaths prevented are 25%, 21%, 18% and 17%, respectively.

Prolonged antibiotic prophylaxis may be indicated if either exposure levels to spores are high, or there is significant risk of re-exposure from either re-aerosolization of spores during remediation of contaminated areas or from another bioterrorist event. Shortening a long course of antibiotics may be critical because of difficulty in maintaining adherence, limited supplies of antibiotics, resistance and adverse events associated with long-term



**Figure 2** Days that antibiotic therapy can be shortened by post-exposure vaccination versus vaccine efficacy. Assumes the probability that persons are protected from disease by vaccine is the vaccine efficacy and that protection begins at or before cessation of antibiotics. Antibiotic efficacy is assumed to be 100%.

antibiotic therapy. We examined the trade-offs associated with shortening the antibiotic regimen by replacing antibiotic protection with post-exposure vaccination. Vaccination confers long-term immunity, whereas antibiotics are effective for only as long as they are being taken. However, that advantage of vaccination must be balanced by the fact that vaccine efficacy is not 100%. We calculated how many days an antibiotic regimen could be shortened by post-exposure vaccination while not increasing total disease risk, assuming antibiotic efficacy was 100%. The results depend on the vaccine efficacy (Fig. 2). For example, if the vaccine efficacy is 95%, 90% and 80%, then antibiotic regimens could be shortened 43, 33 and 23 days, respectively, provided that the level of vaccine efficacy is achieved before antibiotics are terminated. For example, a 120-day antibiotic regimen could be shortened to only 77 days by using a vaccine that was 95% efficacious.

Table 1 suggests that post-exposure antibiotics can prevent at most about 70% of cases. We calculated whether a mass pre-exposure vaccination programme of the general population could significantly increase that prevention rate. Table 3 shows levels of pre-exposure vaccine coverage in the general population that would be necessary to increase the percentages of cases prevented to either 75% or 90%. The results show that the required level of vaccine coverage depends on multiple factors including exposure level, antibiotic response time and adherence; nevertheless, some general conclusions are possible. To prevent 90% of cases, at the very least 63% of the population would need to receive pre-exposure vaccination. That minimal level of vaccine coverage applies when the exposure is low with rapid response and full antibiotic adherence. Generally, vaccine coverage levels that are considerably higher than 63% are required when the response is slower and antibiotic adherence is lower (Table 3).

Table 3 Pre-exposure vaccination to supplement post-exposure intervention

Antibiotic policy		Vaccination policy		
Response time	Adherence	Post-exposure	Pre-exposure coverage* (%)	
			To prevent 75%	To prevent 90%
High exposure				
None	None	No	79	95
None	None	Yes	77	94
Slow	Partial	No	61	88
Slow	Partial	Yes	56	85
Slow	Full	No	56	86
Slow	Full	Yes	56	85
Rapid	Partial	No	43	81
Rapid	Partial	Yes	31	76
Rapid	Full	No	24	73
Rapid	Full	Yes	20	71
Variable moderate exposure				
None	None	No	79	95
None	None	Yes	77	94
Slow	Partial	No	62	88
Slow	Partial	Yes	59	87
Slow	Full	No	57	86
Slow	Full	Yes	56	86
Rapid	Partial	No	44	81
Rapid	Partial	Yes	33	76
Rapid	Full	No	26	74
Rapid	Full	Yes	23	72
Low exposure				
None	None	No	79	95
None	None	Yes	76	94
Slow	Partial	No	55	85
Slow	Partial	Yes	50	83
Slow	Full	No	47	82
Slow	Full	Yes	47	82
Rapid	Partial	No	28	74
Rapid	Partial	Yes	14	69
Rapid	Full	No	4	65
Rapid	Full	Yes	0	63

Calculations use parameter values given in Table 1 footnotes. Antibiotic regimen is 60 days. Vaccine efficacy equals 95%.

\*Percentage of population requiring pre-exposure vaccination necessary to prevent 75% and 90% of cases.

Our analysis shows that minimizing the delay until initiation of antibiotic prophylaxis is key to containing an anthrax outbreak, and reinforces the results of Wein *et al.*<sup>11</sup>. Heightened clinical awareness of symptoms, improved public health surveillance, and efficient systems for mass antibiotic distribution<sup>11,18–19</sup> are all critical components in reducing response time. In the 2001 US outbreak, postal workers did not receive antibiotics until more than 9 days after exposure. However, even if mass antibiotic distribution is completed within 6 days of exposure, at most about 70% of cases could be prevented. Our analysis shows that post-exposure vaccination should not be expected to significantly increase that percentage because of the delays to achieve vaccine protection, but nevertheless can be useful in shortening a prolonged antibiotic regimen. Although mass pre-exposure vaccination programmes of the general population do increase prevention rates, very high vaccine coverage rates are required.

These analyses assumed that there are adequate supplies of antibiotics and vaccine, and that all exposed persons needing post-exposure prophylaxis are completely identified. However, if only a fraction,  $\gamma$ , of exposed persons are identified, then the prevention rates in Table 1 would be overestimated and should then be multiplied by  $\gamma$ . In the event of a large bioterrorism attack,  $\gamma$  could well be less than 1. The analyses assumed that implementation of post-exposure response policies, including identification of exposed persons, does not depend on the numbers of exposed persons. Wein and colleagues<sup>11</sup> considered models that account for the attack size and the queues that would result from limited resources for an emergency response. Developing bioterrorism response policies is challenging because of limited available data on anthrax outbreaks. Although we cannot plan for all possible events of bioterrorism, having public health policy guidelines in place for rapid and rational decision-making are critical elements of public health preparedness. □

## Methods

A probability model based on a competing risks formulation of spore clearance and germination<sup>20</sup> was developed to describe the impact of disease control strategies. The risk that a spore germinates is called  $\lambda$ , and the risk that a spore is naturally cleared from the lung by natural mechanisms—including being expelled through a bronchus, swallowed or killed by macrophages—is called the clearance rate  $\theta$ . These risks are the hazard rates, expressed in units of per day. In addition to these natural clearance mechanisms, a person can be protected from disease by antibiotics or vaccine. If antibiotics are circulating in the body at the time a spore begins to germinate, or if the person has been vaccinated, then it is assumed that the germinating spore is destroyed with probabilities determined by the efficacies of the antibiotics or vaccine. If a spore successfully evades the natural clearance mechanisms, germinates and is not destroyed by vaccine-induced immunity or antibiotics, then it multiplies rapidly, and produces toxin and disease<sup>21–25</sup>.

## Impact of post-exposure prophylaxis

Suppose protection from disease begins at time  $t_1$  then stops at  $t_2$ , and perhaps begins again at  $t_3$ , where the times are measured in days from exposure. Intermittent protection from disease could result from starting and stopping antibiotics, and achieving vaccine-induced immunity.

The probability  $R_1$  that a spore germinates in the interval  $(0, t_1)$  is given by:

$$R_1 = \int_0^{t_1} \lambda e^{-(\lambda+\theta)u} du = \frac{\lambda}{\lambda+\theta} [1 - e^{-(\lambda+\theta)t_1}]$$

The probability  $R_2$  that a spore germinates in the interval  $(t_1, t_2)$  is equal to 0. The probability  $R_3$  that a spore germinates in the interval  $(t_2, t_3)$  is given by:

$$R_3 = \int_{t_2}^{t_3} \lambda e^{-(\lambda+\theta)u} du = \frac{\lambda}{\lambda+\theta} [e^{-(\lambda+\theta)t_2} - e^{-(\lambda+\theta)t_3}]$$

The probability  $R_4$  that a spore germinates after  $t_3$  is equal to 0.

Suppose a person inhales  $D$  spores, and let  $X$  be the number of spores that will ever germinate in that individual. Assuming  $X$  has a Poisson distribution with mean  $\mu = D(R_1 + R_2 + R_3 + R_4)$ , then the cumulative probability that at least one spore germinates is  $c = 1 - e^{-\mu}$ , which is:

$$c = 1 - \exp \left\{ -D \frac{\lambda}{\lambda+\theta} [1 - e^{-(\lambda+\theta)t_1} + e^{-(\lambda+\theta)t_2} - e^{-(\lambda+\theta)t_3}] \right\}$$

The estimates of  $\theta$  from animal data are approximately 0.07 per day, that is, spores are cleared from the lung at about 7% per day<sup>20,26</sup>. That estimate of the clearance rate is consistent with human data on the incubation period from the largest documented human anthrax outbreak in Sverdlovsk<sup>10,20</sup> and observations from the 2001 US outbreak<sup>20</sup>. Estimates of  $\lambda$ , based on estimates of the ID(50) in experimental studies with rhesus



macaque monkeys, suggest that  $\lambda$  ranges between  $3.2 \times 10^{-6}$  and  $8.1 \times 10^{-5}$  per day<sup>27,28</sup>. Thus,  $\lambda$  is much smaller than  $\theta$ . The dose to cause disease with probability  $p = P/100$  is called the infectious dose ID( $P$ ) and is given by<sup>20</sup>  $D = -(\lambda + \theta)\ln(1 - p)/\lambda$ . Substituting this expression for  $D$  into the formula for  $c$  and approximating  $\lambda + \theta \approx \theta$ , we find

$$c = 1 - (1 - p)^{(1 - e^{-\theta t_1} + e^{-\theta t_2} - e^{-\theta t_3})} \quad (1)$$

Thus, we have the surprising mathematical result that it is not necessary to have a precise value for  $\lambda$  to obtain our key results.

We divide the population into subgroups in which the periods of disease protection are the same for all persons in a subgroup. For example, consider a subgroup in which antibiotic protection begins at time  $t_1$  and stops at time  $t_2$ , then we set  $t_3$  to infinity in equation (1). Suppose further that post-exposure vaccine protection is achieved at  $t_v$ , which occurs before antibiotic protection stops ( $t_v < t_2$ ), then  $t_2$  and  $t_3$  are set to infinity. If vaccine protection occurs after antibiotic protection stops then  $t_3$  is set equal to  $t_v$ . The cumulative probability of disease in the population ( $c$ ) is then a weighted average of the values of  $c$  for each subgroup, where the weights are the proportions of people in subgroups. These proportions are determined by the distributions of times in which vaccine immunity is achieved, and the antibiotic start and stop dates (determined by antibiotic adherence rates and efficacy). Among persons exposed to the ID( $P$ ), the percentage of cases prevented with post-exposure interventions is:

$$\left(1 - \frac{c}{p}\right)100 \quad (2)$$

## Impact of pre-exposure vaccination

The impact of adding a pre-exposure vaccination programme is to reduce the cumulative probability of disease from  $c$  to  $c(1 - \beta\phi)$ , in which  $\phi$  is the vaccine efficacy (that is, the probability that the vaccine protects from disease), and  $\beta$  is the vaccine coverage (that is, the fraction of the population that receives pre-exposure vaccination). The fraction  $r$  of all cases prevented by adding a pre-exposure vaccination programme to the post-exposure prophylaxis strategy among persons exposed to the ID( $P$ ), is:

$$r = 1 - \frac{c(1 - \beta\phi)}{p} \quad (3)$$

If we solve equation (3) for  $\beta$  we obtain  $\beta = [1 - \{(p(1 - r))/c\}\phi]^{-1}$ .

The above equation for  $\beta$  was used to calculate the vaccine coverage rates given in Table 3 with  $r = 0.75$  and  $0.90$ .

## Shortening antibiotic regimen by post-exposure vaccine

Consider two situations. In situation 1, antibiotics stop at time  $t_v$ , at which time vaccine protection begins with probability  $\phi$ . In situation 2, antibiotics are continued for a longer period of time, until  $t_2$ , but there is no vaccine protection. We calculate by how much time antibiotics can be shortened ( $t_2 - t_v$ ) in order that the disease risks in the two situations are equal. We set the two probabilities of disease conditional on being disease free at  $t_v$  equal to each other. Solving that equation and approximating  $\lambda + \theta \approx \theta$ , we obtain  $t_2 - t_v = [-\ln(1 - \phi)]/\theta$ ; this was used to produce Fig. 2.

Received 14 September; accepted 4 October 2004; doi:10.1038/nature03087.

1. Inglesby, T. V. *et al.* Anthrax as a biological weapon: medical and public health management. *J. Am. Med. Assoc.* **281**, 1735–1745 (1999).
2. Inglesby, T. V. *et al.* Anthrax as a biological weapon. Updated recommendations for management. *J. Am. Med. Assoc.* **287**, 2236–2252 (2002).
3. Centers for Disease Control. Notice to readers: use of anthrax vaccine in response to terrorism: supplemental recommendations of the Advisory Committee on Immunization Practices. *Morb. Mortal. Wkly Rep.* **51**, 1024–1026 (2002).
4. Jernigan, J. A. *et al.* Bioterrorism-related inhalation anthrax: The first 10 cases reported in the United States. *Emerg. Infect. Dis.* **17**, 933–944 (2011).
5. Joellenbeck, L., Zwanziger, L., Durch, J. & Strom, B. (eds) *The Anthrax Vaccine: Is it Safe? Does it Work?* (National Academy Press, Washington DC, 2002).
6. National Institute of Allergy and Infectious Diseases. National Institute of Allergy and Infectious Diseases biodefense research agenda for CDC category A agents. (<http://www.niaid.nih.gov/publications/bioterrorism.htm>) (2002).
7. Centers for Disease Control. Update: investigation of bioterrorism-related anthrax. *Morb. Mortal. Wkly Rep.* **50**, 1008–1010 (2001).
8. Shepard, C. W. *et al.* Antimicrobial post-exposure prophylaxis for anthrax: Adverse events and adherence. *Emerg. Infect. Dis.* **8**, 1124–1132 (2002).
9. Brookmeyer, R. & Blades, N. Prevention of inhalational anthrax in the U.S. outbreak. *Science* **295**, 1861 (2002).
10. Meselson, M. *et al.* The Sverdlovsk anthrax outbreak of 1979. *Science* **266**, 1202–1208 (1994).
11. Wein, L. M., Craft, D. L. & Kaplan, E. H. Emergency response to an anthrax outbreak. *Proc. Natl Acad. Sci. USA* **100**, 4346–4351 (2003).
12. Committee on Research and Development Needs for Improving Civilian Medical Response to Chemical and Biological Terrorism Incidents. *Chemical and Biological Terrorism* (National Academy Press, Washington DC, 1999).
13. Kaplan, E. H., Craft, D. L. & Wein, L. M. Emergency response to a smallpox attack: the case for mass vaccination. *Proc. Natl Acad. Sci. USA* **99**, 10935–10940 (2002).
14. Bozette, S. *et al.* A model for a smallpox-vaccination policy. *N. Engl. J. Med.* **348**, 416–425 (2003).
15. Halloran, M. E., Longini, I. M. Jr, Nizam, A. & Yang, Y. Containing bioterrorist smallpox. *Science* **298**, 1428–1432 (2002).
16. Ferguson, N. M. *et al.* Planning for smallpox outbreak. *Nature* **425**, 681–685 (2003).
17. Fraser, C., Riley, S., Anderson, R. M. & Ferguson, N. M. Factors that make an infectious disease outbreak controllable. *Proc. Natl Acad. Sci. USA* **101**, 6146–6151 (2004).
18. Kaufmann, A., Meltzer, M. & Schmid, G. The economic impact of a bioterrorist attack: Are prevention and post attack intervention programs justifiable? *Emerg. Infect. Dis.* **3**, 83–94 (1997).

19. Centers for Disease Control and Prevention. Mass antibiotic dispensing: A primer. (<http://www.phppo.cdc.gov/phtn/antibiotic/default.asp>) (June 2004).
20. Brookmeyer, R., Johnson, E. & Bollinger, R. Modeling the optimum duration of antibiotic prophylaxis in an anthrax outbreak. *Proc. Natl Acad. Sci. USA* **100**, 10129–10132 (2003).
21. Brachman, P. S. Inhalation anthrax. *Ann. NY Acad. Sci.* **353**, 83–93 (1980).
22. Ross, J. M. The pathogenesis of anthrax following the administration of spores by the respiratory route. *J. Pathol. Bacteriol.* **73**, 484–494 (1957).
23. Friedlander, A. M. *et al.* Post-exposure prophylaxis against experimental inhalation anthrax. *J. Infect. Dis.* **167**, 1239–1242 (1993).
24. Friedlander, A. M. Anthrax: clinical features, pathogenesis and potential biological warfare threat. *Curr. Clin. Top. Infect. Dis.* **20**, 335–349 (2000).
25. Cieslak, T. J. & Edward, M. E. Clinical and epidemiological principles of anthrax. *Infect. Dis.* **5**, 51–55 (1999).
26. Henderson, D. W., Peakcock, S. & Belton, F. C. Observations on the prophylaxis of experimental pulmonary anthrax in the monkey. *J. Hyg. (Lond.)* **54**, 28–35 (1956).
27. Haas, C. N. On the risk of mortality to primates exposed to anthrax spores. *Risk Anal.* **22**, 189–193 (2002).
28. Watson, A. & Keir, D. Information on which to base assessments of risk from environmental contaminated with anthrax spores. *Epidemiol. Infect.* **113**, 479–490 (1994).

**Acknowledgements** The authors acknowledge the comments of the Anthrax Modeling Working Group of the Secretary's Council on Public Health Preparedness of the Department of Health and Human Services. This research was partially funded by the Fogarty International Center and a grant from the National Institute of Allergy and Infectious Diseases.

**Competing interests statement** The authors declare that they have no competing financial interests.

**Correspondence** and requests for materials should be addressed to R.B. (rbrook@jhsph.edu).

# Transmissibility of 1918 pandemic influenza

Christina E. Mills<sup>1</sup>, James M. Robins<sup>1,2</sup> & Marc Lipsitch<sup>1,3</sup>

<sup>1</sup>Department of Epidemiology, <sup>2</sup>Department of Biostatistics, and <sup>3</sup>Department of Immunology and Infectious Diseases, Harvard School of Public Health, 677 Huntington Avenue, Boston, Massachusetts 02115, USA

The 1918 influenza pandemic killed 20–40 million people worldwide<sup>1</sup>, and is seen as a worst-case scenario for pandemic planning. Like other pandemic influenza strains, the 1918 A/H1N1 strain spread extremely rapidly. A measure of transmissibility and of the stringency of control measures required to stop an epidemic is the reproductive number, which is the number of secondary cases produced by each primary case<sup>2</sup>. Here we obtained an estimate of the reproductive number for 1918 influenza by fitting a deterministic SEIR (susceptible-exposed-infectious-recovered) model to pneumonia and influenza death epidemic curves from 45 US cities: the median value is less than three. The estimated proportion of the population with A/H1N1 immunity before September 1918 implies a median basic reproductive number of less than four. These results strongly suggest that the reproductive number for 1918 pandemic influenza is not large relative to many other infectious diseases<sup>3</sup>. In theory, a similar novel influenza subtype could be controlled. But because influenza is frequently transmitted before a specific diagnosis is possible and there is a dearth of global antiviral and vaccine stores, aggressive transmission reducing measures will probably be required.

The emergence of a new pandemic influenza subtype is expected<sup>3</sup>. Pandemics have occurred regularly throughout this century (1918, 1957 and 1968)<sup>4</sup>, and opportunities for the generation of new subtypes<sup>5</sup> have persisted and probably expanded<sup>3</sup>. Pandemic influenza spreads rapidly, has high attack rates and kills millions of people worldwide<sup>1,4,6</sup>. The 1918 pandemic was particularly destructive. The case fatality proportion (CFP) was ten times higher than in